

Esercizio 0. Funzione generatrice dei momenti di una Normale bidimensionale:

$$\begin{aligned} \mathbb{E}(e^{tZ}) &= \mathbb{E}(e^{t_1X+t_2Y}) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} e^{t_1x+t_2y} f(x,y) dx dy = \\ &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} e^{t_1x+t_2y} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left\{ -\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_x}{\sigma_x}\right)^2 - 2\rho\frac{x-\mu_x}{\sigma_x}\frac{y-\mu_y}{\sigma_y} + \right. \right. \\ &\quad \left. \left. + \left(\frac{y-\mu_y}{\sigma_y}\right)^2 \right] \right\} dx dy = \end{aligned}$$

Applichiamo ora la sostituzione: $u = \frac{x-\mu_x}{\sigma_x}$ e $v = \frac{y-\mu_y}{\sigma_y}$, quindi $x = \sigma_x u + \mu_x$ e $y = \sigma_y v + \mu_y$,
 $du = \frac{dx}{\sigma_x}$ e $dv = \frac{dy}{\sigma_y}$.

$$\begin{aligned} &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{e^{t_1\mu_x+t_2\mu_y}}{2\pi\sqrt{1-\rho^2}} \frac{e^{t_1\sigma_x u+t_2\sigma_y v}}{2\pi\sqrt{1-\rho^2}} \exp\left\{ -\frac{1}{2(1-\rho^2)} \left[u^2 - 2\rho uv + v^2 \right] \right\} dudv = \\ &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{e^{t_1\mu_x+t_2\mu_y}}{2\pi\sqrt{1-\rho^2}} \exp\left\{ -\frac{1}{2(1-\rho^2)} \left[u^2 - 2\rho uv + v^2 - 2(1-\rho^2)t_1\sigma_x u + \right. \right. \\ &\quad \left. \left. - 2(1-\rho^2)t_2\sigma_y v \right] \right\} dudv = \end{aligned}$$

Consideriamo ora solo l'argomento dell'esponenziale dipendente dalle due variabili, completiamo il quadrato rispetto ad u e v aggiungendo e togliendo, quindi, opportune quantità, esso diventa:

$$\begin{aligned} &= \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(u - \rho v - (1-\rho^2)t_1\sigma_x \right)^2 + \right. \right. \\ &\quad \left. \left. - (1-\rho^2)^2 t_1^2 \sigma_x^2 - \rho^2 v^2 - 2(1-\rho^2)\rho t_1\sigma_x v + v^2 - 2(1-\rho^2)t_2\sigma_y v \right] \right\} = \\ &= \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(u - \rho v - (1-\rho^2)t_1\sigma_x \right)^2 + (1-\rho^2) \left(v^2 - (1-\rho^2)t_1^2\sigma_x^2 + \right. \right. \right. \\ &\quad \left. \left. - 2\rho t_1\sigma_x v - 2t_2\sigma_y v \right) \right] \right\} = \\ &= \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(u - \rho v - (1-\rho^2)t_1\sigma_x \right)^2 + (1-\rho^2) \left(v - \rho t_1\sigma_x - t_2\sigma_y \right)^2 + \right. \right. \\ &\quad \left. \left. - (1-\rho^2)(t_1^2\sigma_x^2 + t_2^2\sigma_y^2 + 2\rho t_1 t_2 \sigma_x \sigma_y) \right] \right\} = \end{aligned}$$

Definiamo ora $w = \frac{u - \rho v - (1 - \rho^2)t_1 \sigma_x}{\sqrt{1 - \rho^2}}$ e $z = v - \rho t_1 \sigma_x - t_2 \sigma_y$ e quindi $dw = \frac{du}{\sqrt{1 - \rho^2}}$ e $dz = dv$, l'esponenziale diventa:

$$\begin{aligned} &= \exp\left\{-\frac{1}{2}\left[(w)^2 + (z)^2 - (t_1^2 \sigma_x^2 + t_2^2 \sigma_y^2 + 2\rho t_1 t_2 \sigma_x \sigma_y)\right]\right\} = \\ &= \exp\left\{-\frac{1}{2}\left[(w)^2 + (z)^2\right]\right\} \exp\left\{\frac{1}{2}\left[-(t_1^2 \sigma_x^2 + t_2^2 \sigma_y^2 + 2\rho t_1 t_2 \sigma_x \sigma_y)\right]\right\} \end{aligned}$$

Quindi tornando all'intero integrale:

$$\begin{aligned} \mathbb{E}(e^{tZ}) &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{e^{t_1 \mu_x + t_2 \mu_y} e^{\frac{1}{2}(t_1^2 \sigma_x^2 + t_2^2 \sigma_y^2 + 2\rho t_1 t_2 \sigma_x \sigma_y)}}{2\pi} e^{-\frac{1}{2}(w^2 + z^2)} dw dz \\ &= e^{t_1 \mu_x + t_2 \mu_y + \frac{1}{2}(t_1^2 \sigma_x^2 + t_2^2 \sigma_y^2 + 2\rho t_1 t_2 \sigma_x \sigma_y)} \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{1}{2\pi} e^{-\frac{1}{2}(w^2 + z^2)} dw dz = \\ &= e^{t_1 \mu_x + t_2 \mu_y + \frac{1}{2}(t_1^2 \sigma_x^2 + t_2^2 \sigma_y^2 + 2\rho t_1 t_2 \sigma_x \sigma_y)} \end{aligned}$$

poiché

$$\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{1}{2\pi} e^{-\frac{1}{2}(w^2 + z^2)} dw dz = 1$$

essendo la funzione integranda la densità di una normale bidimensionale che integrata dà quindi 1.

Esercizio 1.

$$\begin{aligned} \text{(a)} \quad \mathbb{P}(\min\{X, Y\} \leq x) &= 1 - \mathbb{P}(\min\{X, Y\} > x) = 1 - \mathbb{P}(X > x, Y > x) = \\ &= (\text{indipendenza}) 1 - \mathbb{P}(X > x)\mathbb{P}(Y > x) = 1 - \mathbb{P}(X > x)^2 = 1 - \left[\sum_{k=x+1}^{+\infty} 2^{-k}\right]^2, \end{aligned}$$

$$\begin{aligned} \text{(b)} \quad \mathbb{P}(Y > X) &= \sum_{k=1}^{+\infty} \mathbb{P}(Y > X, X = k) = \sum_{k=1}^{+\infty} \mathbb{P}(Y > X | X = k)\mathbb{P}(X = k) = \\ &= \sum_{k=1}^{+\infty} \mathbb{P}(Y > k)\mathbb{P}(X = k) = \sum_{k=1}^{+\infty} 2^{-k} 2^{-k} = \sum_{k=1}^{+\infty} 4^{-k} = \sum_{j=1}^{+\infty} \left(\frac{1}{4}\right)^{j+1} = \frac{1}{4} \cdot \frac{4}{3} = \frac{1}{3}, \end{aligned}$$

$$\begin{aligned} \text{(c)} \quad \mathbb{P}(X = Y) &= \sum_{k=1}^{+\infty} \mathbb{P}(X = Y = k) = \sum_{k=1}^{+\infty} \mathbb{P}(X = k)\mathbb{P}(Y = k) = \\ &= \sum_{k=1}^{+\infty} 2^{-k} 2^{-k} = \sum_{k=1}^{+\infty} 4^{-k} = \frac{1}{3}, \end{aligned}$$

$$\begin{aligned} \text{(d)} \quad \mathbb{P}(X \geq kY) &= \sum_{x=1}^{+\infty} \mathbb{P}(X \geq kY, Y = x) = \sum_{x=1}^{+\infty} \mathbb{P}(X \geq kx)\mathbb{P}(Y = x) = \sum_{x=1}^{+\infty} 2^{-kx+1} 2^{-x} = \\ &= 2 \sum_{x=1}^{+\infty} 2^{-x(k+1)} = 2 \sum_{x=1}^{+\infty} \left[\left(\frac{1}{2}\right)^{k+1}\right]^x = 2 \sum_{j=0}^{+\infty} \left[\left(\frac{1}{2}\right)^{k+1}\right]^{j+1} = \\ &= 2 \left(\frac{1}{2}\right)^{k+1} \frac{1}{1 - \left(\frac{1}{2}\right)^{k+1}} = \frac{2}{2^{k+1} - 1}. \end{aligned}$$

Esercizio 2.

$$X_i = \begin{cases} 1 & p \\ 0 & 1-p \end{cases} \quad i = 1, \dots, N, \quad S_N = \sum_{i=1}^N X_i$$

$$\begin{aligned} \mathbb{P}(S_N = k) &= \sum_{n=k}^{+\infty} \mathbb{P}(S_N = k | N = n) \mathbb{P}(N = n) = \sum_{n=k}^{+\infty} \binom{n}{k} p^k (1-p)^{n-k} \frac{\lambda^n e^{-\lambda}}{n!} = \\ &= \frac{(\lambda p)^k e^{-\lambda p}}{k!} \sum_{n=k}^{+\infty} \frac{\lambda^{n-k} (1-p)^{n-k} e^{-\lambda(1-p)}}{(n-k)!} = \frac{(\lambda p)^k e^{-\lambda p}}{k!}, \end{aligned}$$

poiché $\frac{\lambda^{n-k} (1-p)^{n-k} e^{-\lambda(1-p)}}{(n-k)!}$ è la distribuzione di una $\text{Po}(\lambda(1-p))$.

Quella ottenuta è la distribuzione di una $\text{Po}(\lambda p)$.

Esercizio 3.

$$\begin{aligned} \text{(a)} \quad \mathbb{E}(aY + bZ | X) &= \sum_{y,z} (ay + bz) \mathbb{P}(Y = y, Z = z | X = x) = \\ &= a \sum_{y,z} y \mathbb{P}(Y = y, Z = z | X = x) + b \sum_{y,z} z \mathbb{P}(Y = y, Z = z | X = x) = \\ &= a \sum_y y \mathbb{P}(Y = y | X = x) + b \sum_z z \mathbb{P}(Z = z | X = x) = a\mathbb{E}(Y | X) + b\mathbb{E}(Z | X), \\ \text{(b)} \quad \mathbb{E}(X | Y) &= \sum_x x \mathbb{P}(X = x | Y = y) = \sum_x x \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(Y = y)} = \mathbb{E}(X). \end{aligned}$$

Esercizio 4.

$$\begin{aligned} \text{(a)} \quad 1 &= \int_0^{+\infty} dx \int_x^{+\infty} dy cx(y-x)e^{-y} = \int_0^{+\infty} dx cx \left[\int_x^{+\infty} ye^{-y} dy - x \int_x^{+\infty} e^{-y} dy \right] = \\ &= \int_0^{+\infty} dx cx \left[-ye^{-y} \Big|_x^{+\infty} + \int_x^{+\infty} e^{-y} dy - x(-e^{-y}) \Big|_x^{+\infty} \right] = \\ &= \int_0^{+\infty} dx cx \left[xe^{-x} + (-e^{-y}) \Big|_x^{+\infty} - xe^{-x} \right] = \int_0^{+\infty} dx cx e^{-x} = \\ &= c \left[-xe^{-x} \Big|_0^{+\infty} + \int_0^{+\infty} e^{-x} dx \right] = c(-e^{-x}) \Big|_0^{+\infty} = c \Rightarrow c = 1. \end{aligned}$$

$$(b) \quad f_Y(y) = \int_0^y f(x, y) dx = \int_0^y x(y-x)e^{-y} dx = e^{-y} \left[\int_0^y xy \, dx - \int_0^y x^2 \, dx \right] =$$

$$= e^{-y} \left[y \frac{y^2}{2} - \frac{y^3}{3} \right] = \frac{1}{6} e^{-y} y^3.$$

$$f_X(x) = \int_x^{+\infty} x(y-x)e^{-y} dy = x \left[\int_x^{+\infty} ye^{-y} dy - x \int_x^{+\infty} e^{-y} dy \right] =$$

$$= x \left[-ye^{-y} \Big|_x^{+\infty} + \int_x^{+\infty} e^{-y} dy - x(-e^{-y}) \Big|_x^{+\infty} \right] =$$

$$= x \left[xe^{-x} + (-e^{-y}) \Big|_x^{+\infty} - xe^{-x} \right] = xe^{-x}.$$

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{x(y-x)e^{-y}}{\frac{1}{6}e^{-y}y^3} = 6x \frac{y-x}{y^3}$$

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \frac{x(y-x)e^{-y}}{xe^{-x}} = (y-x)e^{x-y}$$

$$(c) \quad \mathbb{E}(X|Y) = \int_0^y x f_{X|Y}(x|y) dx = \int_0^y 6x^2 \frac{y-x}{y^3} dx = \frac{6}{y^3} \left[\int_0^y x^2 y \, dx - \int_0^y x^3 \, dx \right] =$$

$$= \frac{6}{y^3} \left[\frac{y^4}{3} - \frac{y^4}{4} \right] = \frac{6}{y^3} \frac{y^4}{12} = \frac{y}{2} \quad \text{e} \quad \mathbb{E}(X|Y) = \frac{1}{2} Y.$$

$$\mathbb{E}(Y|X) = \int_x^{+\infty} f_{Y|X}(y|x) dy = \int_x^{+\infty} y(y-x)e^{x-y} dy =$$

$$= \int_x^{+\infty} y^2 e^{x-y} dy - x \int_x^{+\infty} ye^{x-y} dy =$$

$$= e^x \left[-y^2 e^{-y} \Big|_x^{+\infty} + \int_x^{+\infty} 2ye^{-y} dy \right] - xe^x \left[-ye^{-y} \Big|_x^{+\infty} + \int_x^{+\infty} e^{-y} dy \right] =$$

$$= e^x \left[x^2 e^{-x} + 2(-ye^{-y} \Big|_x^{+\infty} + \int_x^{+\infty} e^{-y} dy) \right] - xe^x [xe^{-x} - e^{-y} \Big|_x^{+\infty}] =$$

$$= e^x [x^2 e^{-x} + 2xe^{-x} - 2e^{-y} \Big|_x^{+\infty}] - xe^x [xe^{-x} + e^{-x}] = x^2 + 2x + 2 - x^2 - x =$$

$$= x + 2 \cdot \frac{1}{2} \quad \text{e} \quad \mathbb{E}(Y|X) = X + 2.$$

Esercizio 5.

$$\mathbb{E}(e^{tX}) = \sum_{k=0}^{+\infty} e^{tk} \mathbb{P}(X = k) = \sum_{k=0}^{+\infty} e^{tk} \frac{\lambda^k}{k!} e^{-\lambda} = \frac{e^{-\lambda}}{e^{-\lambda e^t}} \sum_{k=0}^{+\infty} \frac{(\lambda e^t)^k}{k!} e^{-\lambda e^t} = e^{-\lambda(1-e^t)}.$$

$$\mathbb{E}(X) = \frac{d}{dt} \mathbb{E}(e^{tX}) \Big|_{t=0} = e^{-\lambda(1-e^t)} (\lambda e^t) \Big|_{t=0} = \lambda.$$

$$\mathbb{E}(X^2) = \frac{d^2}{dt^2} \mathbb{E}(e^{tX}) \Big|_{t=0} = e^{-\lambda(1-e^t)} (\lambda e^t)^2 + e^{-\lambda(1-e^t)} (\lambda e^t) \Big|_{t=0} = \lambda^2 + \lambda.$$

$$\Rightarrow \text{Var}(X) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2 = \lambda^2 + \lambda - \lambda^2 = \lambda.$$

$$\mathbb{E}(e^{tS}) = \mathbb{E}\left(e^{t\sum_{i=1}^n X_i}\right) = \mathbb{E}\left(\prod_{i=1}^n e^{tX_i}\right) = (\mathbb{E}(e^{tX_1}))^n = \left(e^{-\lambda(1-e^t)}\right)^n = e^{-n\lambda(1-e^t)}$$

e questa è la f.g.m. di una $Po(n\lambda)$.

Esercizio 6.

$$\mathbb{E}(e^{tX}) = \int_0^{+\infty} e^{tx} e^{-\lambda x} dx = \int_0^{+\infty} \lambda e^{-(\lambda-t)x} dx = -\lambda \frac{e^{-(\lambda-t)x}}{\lambda-t} \Big|_0^{+\infty} = \frac{\lambda}{\lambda-t}.$$

$$\mathbb{E}(X) = \frac{d}{dt} \mathbb{E}(e^{tX}) \Big|_{t=0} = \frac{\lambda}{(\lambda-t)^2} \Big|_{t=0} = \frac{1}{\lambda}.$$

$$\mathbb{E}(X^2) = \frac{d^2}{dt^2} \mathbb{E}(e^{tX}) \Big|_{t=0} = \frac{2\lambda}{(\lambda-t)^3} \Big|_{t=0} = \frac{2}{\lambda^2}.$$

$$\Rightarrow \text{Var}(X) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}.$$